

RESEARCH ARTICLE

WILEY

Financial conservatism, firm value and international business risk: Evidence from emerging economies around the global financial crisis

Michael Machokoto¹  | Geoffry Areneke²  | Davis Nyangara³

¹Faculty of Business and Law, University of Northampton, Waterside Campus, Northampton, UK

²Department of Accounting, Finance and Banking, Manchester Metropolitan University Business School, Faculty of Business and Law, Manchester Metropolitan University, Manchester, UK

³Department of Finance, National University of Science and Technology, Bulawayo, Zimbabwe

Correspondence

Michael Machokoto, Faculty of Business and Law, University of Northampton, Waterside Campus, Northampton, NN1 5PH, UK.
Email: michael.machokoto@northampton.ac.uk

Funding information

Manchester Metropolitan University

Abstract

The increase in debt-free or under-levered firms (financial conservatism) is one of the most recent stylized puzzles that cannot be explained within the context of extant capital structure theories. In this paper, we exploit the 2008–09 contractions in credit supply in a *quasi-natural experiment* to examine whether financial conservatism affects firm value. Using a large sample of firms from seven African countries over the period 2003–2012, we find strong evidence that financial conservatism mitigates the adverse effect of contractions in credit supply on firm value for both local and international firms. Our results suggest that financial conservatism is an effective strategy for managing risks arising from contractions in credit supply and international business exposure. These findings provide novel empirical evidence on the value relevance of financial conservatism which shields firms from the adverse and far-reaching effects of contractions in credit supply.

KEYWORDS

financial conservatism, firm value, international business risk, cross-listing, global financial crisis, legal origin

1 | INTRODUCTION

One of the most recent stylized puzzles in corporate finance is the increase in debt-free or under-levered (financial conservatism, thereon) firms which is contrary to predictions of the main capital structure theories. The trend implies that firms are increasingly forgoing the benefits associated with debt finance such as interest tax shield, lower costs of capital and the disciplinary role of debt (Bessler, Drobetz, Haller, & Meier, 2013; Bigelli,

Martín-Ugedo, & Sánchez-Vidal, 2014; Graham & Harvey, 2001). This puzzling phenomenon has motivated several studies on the determinants of financial conservatism. For example, Bigelli et al. (2014) show that financial conservatism is more pronounced in smaller firms and firms with low intangible assets and effective tax rates. Bessler et al. (2013) also report that financial conservatism is higher in common law countries, countries with higher creditor protection and in jurisdictions with dividend relief tax systems. Notwithstanding the contributions of these papers, it remains unclear whether financial conservatism has any financial benefit or cost, especially during periods of heightened uncertainty and

[Correction added on 22 September 2020, after first online publication: The funding information has been updated in this current version.]

This is an open access article under the terms of the Creative Commons Attribution-NonCommercial License, which permits use, distribution and reproduction in any medium, provided the original work is properly cited and is not used for commercial purposes.

© 2020 The Authors. *International Journal of Finance & Economics* published by John Wiley & Sons Ltd.

contractions in credit supply such as the global financial crisis. In this paper, we address the above research gap and provide new insights on whether financial conservatism enhances or preserves firm value during the 2008–09 credit supply shock within the context of emerging markets.

The effects of the global financial crisis (GFC) are still being felt, and much is yet to be understood about how and why it affected some economic agents more than others at firm, industry and country level (Anginer & Demirguc-Kunt, 2014; Dawood, Horsewood, & Strobel, 2017).¹ Our study sheds light on the above research question and is important for two main reasons. First, as the link between financing and real decisions is contentious, the 2008–09 exogenous credit supply shock provides a rare and unique *quasi-natural experiment* setting that enables us to examine this issue directly.² Second, the 2008–09 contraction in credit supply offers an opportunity to examine how economic shocks originating in developed countries propagate and affect firms in less integrated and understudied emerging economies. This is of particular importance as there is a general consensus that the effects of the GFC vary in both intensity and duration with the level of economic integration (Allen & Giovannetti, 2011; Berman & Martin, 2012; Duygun, Isaksson, & Kaulich, 2016). By focusing on the GFC, we provide direct empirical evidence on the adverse effects of the crisis across emerging economies and whether these are moderated by financing structures.

Using a large sample of publicly listed African firms, we examine the effects of financial conservatism on firm value over the period 2003–2012. We contend that as the GFC was an unexpected credit supply shock, firms had limited or no time to adjust their financing policies, hence, our estimates are less likely to be biased by reverse causality. As expected, we find that the GFC had a significant negative effect on firm value. On average, firm value decreased by 14%–20% after 2008. This decrease is robust to different definitions of firm value, the window-period used and other factors that are known to affect firm value. However, in our main analysis we find that this negative effect is insignificant and less pronounced for firms that are financially conservative. This finding is novel as the prior literature on financial conservatism is rather confined to the examination of the determinants of conservatism and does not investigate when and where it is most beneficial.³ Thus, our results suggest that financial conservatism has several important implications on firm value in the aftermath of the GFC as it increases resilience to contractions in credit supply.

We further find that the severity of the 2008–09 contractions in credit supply varies across firms and countries as well as over time. Our results show that the panic occurred earlier in Africa relative to developed countries, with the decrease in firm value peaking in 2006, and diminishing thereafter. This is inconsistent with the US studies of Almeida and

Campello (2007) and Chava and Purnanandam (2011) that report severe adverse effects during the 2008–09 period. We further find that firms with international exposure, through either cross-listing or foreign sales, were more affected than purely domestic firms and that this effect is moderated by financial conservatism. This implies that firms with high exposure to international markets can mitigate the associated risks by adopting conservative financing policies. Our results further show that firms in common law countries (Nigeria, South Africa, Egypt and Kenya), which are more integrated with the US and the UK (where the crisis originated), were more affected by the crisis than those in civil law countries (Morocco, Tunisia and Ivory Coast). Hence, high levels of economic integration and similar legal systems expose firms in vulnerable countries to severe credit supply shocks originating from developed economies.

We contribute to the literature in five ways. First, we show that financial conservatism is beneficial during the 2008–09 contractions in credit supply. This implies that firms with conservative financial policies, specifically those that are debt-free or unlevered, are better positioned to manage the adverse effect of credit supply shocks relative to those with less conservative financial policies. Second, we contend that, as market imperfections are more apparent in emerging economies, this should result in a pronounced effect of financing decisions on firm value. In line with this prediction, we find that the adverse effects of the GFC increase with leverage since financially conservative firms in our context were less affected relative to non-conservative ones. Third, our results show that economic ties and international business exposure increase vulnerability to economic shocks, but this effect diminishes with financial conservatism. Fourth, we find that firms in common law countries that have closer links to the US and the UK were more adversely affected by the GFC, but this is less visible for financially conservative firms. Finally, we show that firms in African countries that are less integrated and understudied were also adversely affected by the GFC.⁴

The remainder of the paper proceeds as follows. Section 2 presents the literature review and hypotheses. In Section 3, we present the methodology used. Section 4 explains the data and variable construction. Section 5 presents and discusses the empirical results. Section 6 presents the robustness tests. Finally, Section 7 concludes.

2 | LITERATURE AND HYPOTHESES

Our study builds upon four strands of the literature, namely, (1) the effects of contractions in credit supply on corporate decisions, (2) the nexus between financing and real activities, (3) the link between internationalization

and corporate outcomes and (4) the effect of legal origin on corporate outcomes.

2.1 | The GFC, financial conservatism and firm value

Over the past few decades, the frequency and severity of economic crises have increased. The majority of these crises have either been country-specific (e.g., Turkish, Argentinian and Russian crisis of 2001, 1992–2002 and 2014, respectively) or region-specific (e.g., the Asian financial crisis of the late 1990s and the 2010 euro-zone crisis). The exception is the 2008–09 GFC. As the GFC originated from the US subprime mortgage market and propagated to other countries (unlike other previous regional or country-specific crises) it is not surprising that it has motivated considerable research. However, studies on the effects of the GFC have reported mixed results, with Campello, Graham, and Harvey (2010), Chava and Purnanandam (2011), Almeida, Murillo, Bruno, and Scott (2012), Udenio, Hoberg, and Fransoo (2018) and Rehman, Chaudhry, and Hussain (2019) documenting significant adverse effects on corporate decisions, while Chari, Christiano, and Kehoe (2008), Lemmon and Roberts (2010), Kahle and Stulz (2013), Caporale, Lodh, and Nandy (2018), Zouaghi, Sánchez, and Martínez (2018) and Choudhry, Hassan, and Shabi (2020) report no or low effects.⁵ The findings of the latter studies are contrary to the central prediction that firms with weaker balance sheets are more likely to be credit rationed during periods of heightened uncertainty (Holmstrom & Tirole, 1997). In addition to this lack of consensus in the literature, there is a dearth of studies that have examined whether firm financial policies provide a shield or increase exposure to the 2008–09 contractions in credit supply.

We address the above research gap and contribute to the literature by examining whether conservative financial policies were important in protecting or shielding firms from the adverse effects of the GFC. Our focus is motivated by the growing interest among academics in understanding why firms are increasingly adopting conservative financial policies (zero and low leveraging) (see Bessler et al., 2013; Bigelli et al., 2014; Dang, 2013; Devos et al., 2012). By adopting such policies, firms are implicitly choosing to forgo the benefits of debt finance, which is inconsistent with the mainstream theories. While extant studies have investigated the determinants of financial conservatism, the question of whether financial conservatism is value relevant during contractions in credit supply is yet to be examined. We contend that unlevered or conservative firms are less likely to be affected by contractions in credit supply since they are more reliant on internal and equity financing.

Accordingly, they ought to be better positioned to manage the adverse effects of contractions in credit supply relative to levered firms as they have prior experience in managing investments without debt financing. This is in line with economic theories of predation in which firms may voluntarily decide to lower their leverage to gain a competitive edge and market share during economic slumps (see Chevalier & Scharfstein, 1996). Against this background, we, therefore, formulate and test the following hypothesis.

Hypothesis: 1 *Financial conservatism enhances firm value during contractions in credit supply.*

2.2 | International exposure, legal origin, financial conservatism and firm value

Globalization has spurred the movement of firms across national boundaries to increase market share, profitability and competitive advantage. This has motivated recent research focusing on the interface between international business and finance (see Aggarwal, Erel, Stulz, & Williamson, 2009; Fainshmidt, Judge, Aguilera, & Smith, 2018; Areneke & Kimani, 2019). Several extant studies document that internationalization has several important benefits to the firm, such as improvements in corporate practices, profitability and firm value. For example, Juasrikul, Sahaym, Yim, and Liu (2018) show that international alliances with developed market MNEs (DMMNEs) enhance the firm value of emerging market multinationals (EMMNEs). Similarly, Hsu, Chen, and Cheng (2013) find that international expansion enhances growth opportunities and access to cutting-edge technology. However, not much has been done on investigating whether firm internationalization increases contagion risks and exposure to international shocks, and whether financial policies moderate this effect.

We address this lack of research by focusing on the differential effects of the GFC on cross-listed firms and firms with foreign sales. Firms can achieve internationalization through cross listing in other countries. Drawing on the bonding hypothesis of Coffee (2002), cross listing enhances access to external finance and scrutiny as firms bond with international practices. This improves the competitive edge and the relative performance of cross-listed firms over domestic firms. As a result, cross listing enhances firm value (Charitou & Louca, 2009; Coffee, 2002) and the quality of corporate governance, thus improves access to external finance and mitigates agency costs and information asymmetry problems (see Areneke & Kimani, 2019; Lel & Miller, 2008). However, during contractions in credit supply, cross listing is likely to increase susceptibility to international shocks. We contend that financing policies can

moderate this adverse effect since cross-listed firms with conservative financial policies are less likely to be affected by credit supply shocks. This arises because, unlike levered firms, conservative firms are not dependent on debt financing for their survival. Thus, conservative financial policies reduce or moderate the adverse impact of the contractions in credit supply for cross-listed firms during the GFC. We, therefore, hypothesize the following:

Hypothesis: 2 *Financial conservatism mitigates the adverse effects of the GFC on the value of firms with international business exposure through cross listing.*

As not all internationalised firms are cross-listed, we use foreign sales as another proxy for international business exposure. We posit that during contractions in credit supply, firms with foreign sales are more likely to be affected by contagion effects as their sales shrink both domestically and internationally. However, if our first hypothesis (H1) holds, this severe adverse effect is likely to be moderated by financial conservatism. Financial conservatism, in this case, increases strategic flexibility, which frees the firm to pursue other growth prospects or new markets rather than focus entirely on dealing with debt-overhang issues (which affect levered firms) triggered by the contraction in credit supply. Against this background, we formulate and test the following hypothesis:

Hypothesis: 3 *Financial conservatism mitigates the adverse effects of the GFC on the value of firms with foreign sales.*

Finally, the legal origin of countries ensures strong ties in policies and governance systems such as property rights and investors protection (La Porta, Lopez-De-Silanes, Shleifer, 2006; La Porta, Lopez-De-Silanes, Shleifer, & Vishny, 1997). As noted earlier, Bessler et al. (2013) show that countries with high creditor protection and common law systems have the highest number of financially conservative firms. Drawing on this finding, we conjecture that legal origin ensures complementariness and strong ties with countries that have similar systems. While these ties are beneficial, they also constitute another channel that can aid in the transmission of economic shocks to other countries. As the 2008–09 credit supply shock originated in the US and the UK, we expect firms in common law countries to have been more affected relative to those in civil law systems. Accordingly, we test whether financial conservatism (if H1 holds) is more value relevant to firms in countries that have stronger ties with the US and the UK (common law countries) where the credit supply shock originated. We, thus, hypothesize the following:

Hypothesis: 4 *The mitigating role of financial conservatism on firm value during the GFC is more pronounced for firms in countries with common law legal systems.*

3 | METHODOLOGY

To examine the effects of financial conservatism on firm value, we estimate the following model:

$$y_{ijt} = \alpha + \gamma_1 \text{Crisis}_t + \gamma_2 \text{Crisis}_t \times \text{FC}_i + \beta \mathbf{X}_{ijt-1} + \eta_j + \eta_t + \epsilon_{ijt} \quad (1)$$

where y_{ijt} is the value of firm i in country j at time t ; γ_1 , γ_2 and β are coefficients to be estimated; Crisis_t is a dummy that takes the value of one over the period 2008–2012 and zero otherwise; FC_i is a dummy that takes the value of one for financially conservative firms and zero otherwise; \mathbf{X}_{ijt-1} is a vector of firm-specific characteristics explained below; η_j and η_t are the country and time-fixed effects; and ϵ_{ijt} is the error term. As financial conservatism (FC_i) is time-invariant, the coefficient of FC_i in Equation (1) is absorbed by the firm-fixed effect. For robustness, we use several definitions of financial conservatism (FC_i), namely: zero leverage (ZL) for firms without debt, almost zero leverage (AZL) for firms with less than 1% debt, and ultra-low leverage (UL) for firms with less than 5% debt. Our measures of financial conservatism are based on the pre-crisis period to avoid the possible compounding effects of the credit shock on financing decisions (see Angrist & Pischke, 2008). The vector \mathbf{X}_{ijt-1} consists of return on assets (ROA), sales growth (SG), cash and cash equivalents (Cash), property, plant and equipment (PPE), size (Size) and the median of the dependent variable.⁶

To investigate the variations of the effect of leverage on firm value, we next estimate a modified version of Equation (1) that includes three leverage dummies ($D2_{ijct} - D4_{ijct}$) as follows:-

$$y_{ijt} = \alpha + \sum_{k=2}^4 \lambda_k \mathbf{Dk}_{ijct} + \gamma_1 \text{Crisis}_t + \sum_{k=2}^4 \lambda'_k \mathbf{Dk}_{ijct} \times \text{Crisis}_t + \beta \mathbf{X}_{ijt-1} + \eta_j + \eta_t + \epsilon_{ijt} \quad (2)$$

where $\sum_{k=2}^4 \beta_k \mathbf{Dk}_{ijct}$ is a vector of financing (leverage) policies defined below with slope coefficients from λ_2 to λ_4 . $D1_{ijt}$ (low leverage), $D2_{ijt}$ (low-medium leverage), $D3_{ijt}$ (medium-high leverage) and $D4_{ijt}$ (high leverage) are

TABLE 1 Variable definitions

Variable	Source	Definition
Q	<i>DataStream</i>	Market value of equity plus total debt-to-total assets (firm value [Tobin's q]).
MVE	<i>DataStream</i>	Market value of equity-to-book equity.
TSR	<i>DataStream</i>	Total shareholder return.
ZL	<i>DataStream</i>	Is a dummy variable that takes the value of one if a firm has no debt, and zero otherwise.
AZL	<i>DataStream</i>	Is a dummy variable that takes the value of one if a firm has almost zero leverage ($\leq 1\%$), and zero otherwise.
UL	<i>DataStream</i>	Is a dummy variable that takes the value of one if a firm has less than 5% debt and zero otherwise.
FC	<i>DataStream</i>	A dummy that takes the value of one for financially conservative firms and zero otherwise.
NFC	<i>DataStream</i>	A dummy that takes the value of one for non-financially conservative firms and zero otherwise.
D1	<i>DataStream</i>	Is a dummy variable that takes the value of one if a firm is categorized into the first quantile of the leverage level and zero otherwise.
D2	<i>DataStream</i>	Is a dummy variable that takes the value of one if a firm is categorized into the second quantile of the leverage level and zero otherwise.
D3	<i>DataStream</i>	Is a dummy variable that takes the value of one if a firm is categorized into the third quantile of the leverage level and zero otherwise.
D4	<i>DataStream</i>	Is a dummy variable that takes the value of one if a firm is categorized into the fourth quantile of the leverage level and zero otherwise.
Debt	<i>DataStream</i>	Total debt-to-total assets.
ROA	<i>DataStream</i>	Operating income-to-total assets.
SG	<i>DataStream</i>	Sales growth.
Cash	<i>DataStream</i>	Cash and equivalent-to-total assets.
Size	<i>DataStream</i>	Log of total assets.
Foreign sales	<i>DataStream</i>	Foreign sales-to-total assets.
Dual listing	<i>DataStream</i>	Is a dummy variable that takes the value of one if a firm is dual-listed and zero otherwise.
IndMedQ	<i>DataStream</i>	The industrial median firm value in each year.
IndMedMVE	<i>DataStream</i>	The industrial median MVE in each year.
IndMedTSR	<i>DataStream</i>	The industrial median TSR in each year.
Civil		Country classifications into civil and common law are based on La Porta et al. (1997).

Note: The table lists the definitions and sources of all variables used.

dummy variables that take the value of one and otherwise zero if firm i in industry j at time t is categorized as following one of the four leverage policies. For example, $D1_{ijt}$ is equal to one if a firm is categorized in the first quantile of the leverage level and otherwise zero. We drop $D1_{ijt}$ (low leverage) to avoid the dummy variable trap. To deal with unobserved heterogeneity, we include firm, country and time-fixed effects in all regressions and report standard errors that are clustered at firm level.

4 | DATA

Our sample consists of publicly listed firms extracted from *Thomson Reuters Datastream* over the period 2003–2012. As is standard in the literature, we exclude firms in the financial and utility sectors, those with missing data on key variables and firms with more than 100% growth in assets or sales (see Baek, Kang, & Suh Park, 2004; Enikolopov, Petrova, & Stepanov, 2014). We only retain firms that have at least one observation in

both the pre-crisis (2003–2007) and the crisis periods (2008–2012). All variables used are winsorized at the lower and upper one percentile. Our final unbalanced sample consists of 5,320 firm-year observations (901 firms) from seven African countries (Egypt, Ivory Coast, Kenya, Morocco, Nigeria, South Africa and Tunisia). All variables used are defined in Table 1.

Table 2 presents the summary statistics for all the variables. Panel A, for the main variables, shows an overall mean (*SD*) of 1.578 (0.818), 2.248 (1.723) and 0.226 (0.0574) for firm value (*Q*), market value-to-equity (*MVE*) and total shareholder return (*TSR*), respectively. These summary statistics are in line with the literature.⁷ Comparisons of the periods 2003–2007 and 2008–2012 suggest that the GFC had a significant negative effect on firm value (*Q*), *MVE* and *TSR*. The changes for the other firm-specific factors around the GFC period are not significant and less pronounced, except for *ROA* and size (*Size*). Panel B presents the difference in firm value between the period 2003–2007 and 2008–2012, and across non-financially conservative (*Non-FC*) and financially conservative (*ZL*, *AZL* and *UL*) firms. The results show that the value of financially conservative firms decreased less than that for non-financially conservative firms. This significant and more pronounced decrease in value for non-financially conservative firms (*Non-FCs*) is consistent across different measures of firm value, namely *MVE* and *TSR*. These differences are consistent with our main hypothesis (*H1*) and show *a priori* that conservative financial policies reduce the adverse effects of the GFC on firm value.

Figure 1 plots firm value for the four quantiles of firms categorized by leverage. Figure 1(a) shows that firm value decreases with leverage and, more importantly, it is higher in the pre-crisis period (2003–2007) than in the crisis-period (2008–2012) across all leverage quantiles. In Figure 1(b), we observe similar variations in firm value around the GFC for the box plots based on a shorter window-period (2006–2009). This provides *prima facie* evidence that the GFC had a negative effect on firm value and that this effect varies with the level of leverage.

Figure 2 presents the cross-country differences in firm value. Of particular interest are the variations in firm value between the periods 2003–2007 and 2008–2012, which show that the value of firms in civil law countries (Ivory Coast, Tunisia and Morocco) increased, while that of firms in common law countries decreased. We link this result to differences in the level of integration with the countries where the crisis originated (the US and UK). Thus, firms in countries that have closer economic, legal and social ties to the US and the UK (common law countries) were more affected than those in civil law countries.

Table 3 presents the pairwise correlations of all variables used. Firm value (*Q*) is positively correlated with cash flow, debt, *MVE* and size, while it is negatively correlated with cash. These correlations are as expected, except for cash, which suggests that firms reduce cash holdings to finance investments.

5 | RESULTS

To test our main hypothesis (*H1*), we estimate Equation (1) that relates firm value to financial conservatism (*ZL*, *AZL* and *UL*), a financial crisis dummy (*Crisis*), an interaction term of financial conservatism with the financial crisis and other control variables. Table 4 summarizes our main findings.

Columns (1)–(8) of Table 4 consistently show that the GFC (*Crisis*) has a negative and significant effect on firm value. This result holds for both a longer (2003–2012) and shorter (2006–2009) window-period around the GFC, and is broadly consistent with our first hypothesis (*H1*) and the literature on advanced economies (see Baek et al., 2004; Enikolopov et al., 2014; Gupta, Krishnamurti, & Tourani-Rad, 2013). We further find that the interaction term, *Crisis*×*FC*, is positive and significant, which implies that financially conservative firms were less affected by the credit supply shock. Our results show that the value of levered firms decreased by 0.567 and that for zero-leverage (*ZL*) firms increased by 0.161 ($\gamma_1 + \gamma_2$) (in Column 2). We find similar but lower decreases of 0.155 and 0.309 for the other measures of financial conservatism, namely *AZL* firms in Column (3) and *UL* firms in Column (4), respectively. Columns (5)–(8), for the shorter window-period (2006–2009), show similar results, with *ZL* firms experiencing a moderate increase in firm value. Similarly, the value of *AZL* and *UL* firms decreased by a small and insignificant margin.

Figure 3 plots the estimated coefficients of Equation (1) with indicator variables for 2003–2007, 2009, 2010, 2011 and 2012 interacted with the dummy for non-financially conservative firms (*NFC_t*). These plots indicate that levered firms (non-financially conservative firms (*Non-FC*)) experienced significant decreases in firm value from 2007 onwards. The decrease in firm value that we document is robust to alternative definitions of financial conservatism (as shown in Figure 3(a) and (b)). These results are in line with Table 4 and suggest that financial conservatism enhances or preserves firm value during credit supply shocks.

We next examine the effects of the GFC across firms with different levels of debt financing (leverage). To do this, we estimate Equation (2) augmented with three dummies that capture differences in the level of debt

TABLE 2 Basic statistics

Panel A: Main variables												
Variables	2003–2012			2003–2007			2008–2012			Differences		Kruskal–Wallis
	Mean	Median	SD	Mean	Median	SD	Mean	Median	SD	Mean	Median	
Q	1.578	1.351	0.818	1.707	1.508	0.811	1.479	1.237	0.810	−0.228 ^a	−0.271 ^a	191.210 ^a
MVE	2.248	1.728	1.723	2.566	2.043	1.767	2.002	1.475	1.647	−0.564 ^a	−0.568 ^a	225.758 ^a
TSR	0.226	0.106	0.574	0.421	0.292	0.624	0.075	0.000	0.481	−0.346 ^a	−0.292 ^a	503.887 ^a
TDA	0.157	0.136	0.138	0.157	0.136	0.133	0.158	0.134	0.142	0.001	−0.002	0.769
ROA	0.168	0.165	0.134	0.178	0.174	0.147	0.161	0.158	0.123	−0.017 ^a	−0.016 ^a	34.135 ^a
SG	0.104	0.086	0.198	0.107	0.089	0.186	0.102	0.085	0.206	−0.005	−0.004	0.224
Cash	0.134	0.102	0.116	0.136	0.102	0.118	0.132	0.102	0.115	−0.004	0.000	1.114
Size	14.818	14.951	2.094	14.918	15.128	2.018	14.745	14.731	2.146	−0.173 ^b	−0.397 ^a	7.743 ^b
MedQ	1.418	1.360	0.363	1.394	1.360	0.293	1.436	1.353	0.407	0.042 ^a	−0.007	0.010
MedMVE	1.873	1.733	0.814	1.849	1.764	0.737	1.892	1.685	0.867	0.043 ^d	−0.079	0.158
MedTSR	0.172	0.164	0.277	0.289	0.319	0.244	0.084	0.079	0.267	−0.205 ^a	−0.240 ^a	738.951 ^a
N	5,230			2,279			2,951					
Firms	901											
Panel B: Non-financially and financially conservative firms												
#	Period	Q			MVE			TSR			UL	
		Non-FC	ZL	AZL	UL	Non-FC	ZL	AZL	UL	Non-FC		ZL
(1)	All	1.518	2.049	1.885	1.813	2.171	2.638	2.429	2.543	0.214	0.296	0.307
(2)	2003–2007	1.659	2.039	1.868	1.852	2.547	2.744	2.476	2.623	0.430	0.376	0.410
(3)	2008–2012	1.418	2.058	1.905	1.771	1.906	2.537	2.376	2.457	0.062	0.219	0.190
N		4,157	335	580	1,073	4,157	335	580	1,073	4,157	335	580
Differences (3)–(2)												
Diff-mean		−0.241 ^a	0.019	0.037	−0.081	−0.641 ^a	−0.207	−0.100	−0.166	−0.368 ^a	−0.157 ^b	−0.256 ^a
Diff-median		−0.271 ^a	−0.153	−0.169	−0.177 ^b	−0.589 ^a	−0.022	−0.073	−0.344 ^c	−0.285 ^a	−0.035	−0.210 ^a
Kolmogorov–Smirnov		0.195 ^a	0.116	0.104 ^d	0.140 ^a	0.213 ^a	0.131	0.120 ^c	0.148 ^a	0.280 ^a	0.179 ^b	0.248 ^a
Kruskal–Wallis		176.559 ^a	0.628	0.918	8.214 ^b	212.116 ^a	1.527	1.851	11.766 ^a	422.640 ^a	5.162 ^c	22.230 ^a
												68.352 ^a

Note: The table presents the summary statistics for all variables used. The sample consists of listed non-utility and non-financial firms in selected African countries drawn from *Datastream* over the period 2003–2012. All variables used are defined in Table 1 and are winsorized at the lower and upper one percentiles. a, b, c and d indicate significance at the 0.01%, 1%, 5% and 10% levels, respectively.

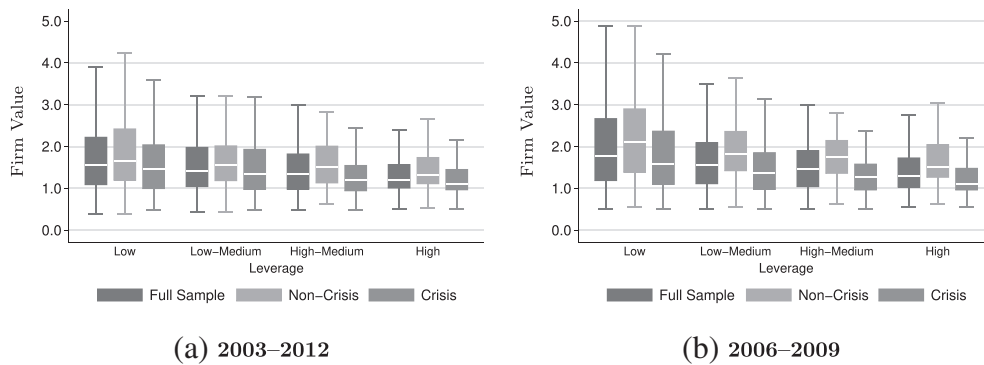


FIGURE 1 Firm value and leverage before and during the crisis. The figure plots firm value for the four subgroups of firms over the sample periods 2003–2012 and 2006–2009. In each year, we partition the firms into four quantiles based on leverage, namely, Low, Low-Medium, High-Medium and High. The central horizontal line in the box is the median while the top and bottom of the box are the 25th and 75th percentiles, respectively. The upper and lower whiskers are the maximum and minimum values, respectively. The sample consists of listed non-utility and non-financial firms in selected African countries drawn from *Datastream* over the periods 2003–2012 and 2006–2009. All variables used are defined in Table 1 and are winsorized at the lower and upper one percentiles

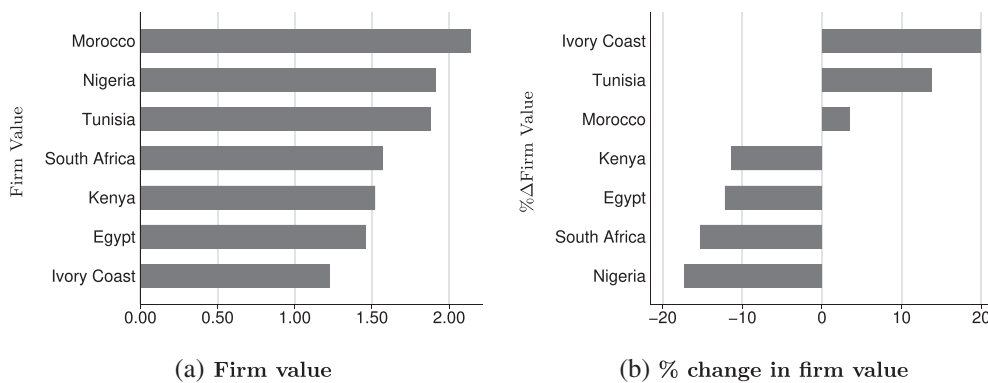


FIGURE 2 Firm value by country before and during the crisis. The figure plots firm value by country (a) and for the percentages change in firm value from the pre-crisis (2003–2007) and crisis (2008–2012) (b). The sample consists of listed non-utility and non-financial firms in selected African countries drawn from *Datastream* over the period 2003–2012. All variables used are defined in Table 1 and are winsorized at the lower and upper one percentiles

financing used by the sample firms. Table 5 summarizes the results for our augmented model.

Consistent with our predictions, the coefficient estimates of the three leverage policies ($D2$, $D3$, $D4$) in Column (1) are negative and insignificant. This suggests that the firms have similar values before the crisis despite the differences in financing structures. However, Column (2) shows that the estimate of the coefficients on the crisis dummy (*Crisis*) and the interaction terms ($Crisis \times D2$, $Crisis \times D3$ and $Crisis \times D4$) are significant. Columns (3) and (4) for the shorter window-period around the GFC (2006–2009) show similar results. This provides further supporting evidence for our first hypothesis (H1) and suggests that firms with high levels of debt experienced a relatively more pronounced decrease in firm value during the GFC.

We next examine the effects of international business exposure on the relationship between firm value and financial conservatism (H2 and H3). We partition firms into two groups based on whether a firm is dual-listed (Yes) or not (No) and whether it has foreign sales (Yes) or not (No). Table 6 summarizes the estimation results for these sub-samples.

Panel A of Table 6 shows that dual-listed firms were more affected by the GFC than domestic or local firms, which suggests negative spillover effects of bonding. However, this effect is less pronounced for financially conservative firms. This implies that financial conservatism moderates or reduces the adverse effects of credit supply shocks on firm value for cross-listed firms. Our results suggest that financial conservatism is a mechanism through which firms with foreign listings can

TABLE 3 Correlation

#	Variables	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	(9)	(10)	(11)	(12)	(13)	(14)
(1)	Q	1													
(2)	MVE	0.879 ^a	1												
(3)	TSR	0.284 ^a	0.279 ^a	1											
(4)	ZL	0.158 ^a	0.066 ^a	0.036 ^c	1										
(5)	AZL	0.139 ^a	0.043 ^b	0.052 ^a	0.730 ^a	1									
(6)	UL	0.153 ^a	0.096 ^a	0.038 ^b	0.503 ^a	0.689 ^a	1								
(7)	TDA	-0.198 ^a	-0.120 ^a	-0.017	-0.289 ^a	-0.393 ^a	-0.496 ^a	1							
(8)	ROA	0.386 ^a	0.281 ^a	0.061 ^a	0.102 ^a	0.091 ^a	0.113 ^a	-0.268 ^a	1						
(9)	SG	0.108 ^a	0.102 ^a	0.033 ^c	-0.035 ^c	-0.024	-0.029 ^d	-0.027 ^d	0.230 ^a	1					
(10)	Cash	0.210 ^a	0.136 ^a	0.036 ^c	0.271 ^a	0.347 ^a	0.315 ^a	-0.384 ^a	0.234 ^a	0.001	1				
(11)	Size	0.093 ^a	0.141 ^a	-0.072 ^a	-0.211 ^a	-0.196 ^a	-0.137 ^a	0.090 ^a	0.055 ^a	0.082 ^a	-0.212 ^a	1			
(12)	MedQ	0.355 ^a	0.347 ^a	-0.074 ^a	0.135 ^a	0.073 ^a	0.136 ^a	-0.126 ^a	0.164 ^a	0.082 ^a	0.072 ^a	0.032 ^c	1		
(13)	MedMVE	0.335 ^a	0.350 ^a	-0.067 ^a	0.112 ^a	0.055 ^a	0.123 ^a	-0.107 ^a	0.178 ^a	0.115 ^a	0.062 ^a	0.061 ^a	0.953 ^a	1	
(14)	MedTSR	0.215 ^a	0.234 ^a	0.129 ^a	0.019	0.039 ^b	0.058 ^a	-0.070 ^a	0.139 ^a	0.073 ^a	0.082 ^a	0.008	0.417 ^a	0.399 ^a	1

Note: The table presents the pairwise correlations for all variables used. The sample consists of listed non-financial and non-utility firms in selected African countries drawn from *Datastream* over the period 2003–2012. All variables used are defined in Table 1 and are winsorized at the lower and upper one percentiles. a, b, c and d indicate significance at the 0.01%, 1%, 5% and 10% levels, respectively.

TABLE 4 The effects of financial conservatism on firm value

Variables	2003–2012				2006–2009			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Constant	2.296 ^a (0.591)	2.130 ^a (0.566)	2.133 ^a (0.566)	2.015 ^a (0.583)	10.808 ^a (1.084)	10.125 ^a (1.050)	10.386 ^a (1.068)	10.290 ^a (1.080)
Crisis	−0.513 ^a (0.038)	−0.567 ^a (0.039)	−0.573 ^a (0.038)	−0.585 ^a (0.039)	−0.431 ^a (0.043)	−0.495 ^a (0.039)	−0.493 ^a (0.041)	−0.493 ^a (0.044)
Crisis×ZL		0.728 ^a (0.152)				0.608 ^a (0.141)		
Crisis×AZL			0.418 ^a (0.097)				0.352 ^a (0.103)	
Crisis×UL				0.276 ^a (0.069)				0.203 ^b (0.076)
ROA	0.979 ^a (0.084)	0.925 ^a (0.077)	0.923 ^a (0.080)	0.959 ^a (0.080)	0.434 ^c (0.202)	0.278 (0.198)	0.316 (0.207)	0.343 (0.213)
SG	0.131 ^a (0.039)	0.121 ^a (0.036)	0.131 ^a (0.037)	0.136 ^a (0.038)	0.119 (0.089)	0.096 (0.087)	0.121 (0.088)	0.129 (0.089)
Cash	1.006 ^a (0.144)	0.944 ^a (0.154)	0.982 ^a (0.151)	1.012 ^a (0.149)	1.371 ^a (0.265)	1.331 ^a (0.264)	1.398 ^a (0.265)	1.422 ^a (0.265)
Size	−0.083 ^c (0.039)	−0.062 ^d (0.037)	−0.067 ^d (0.037)	−0.060 (0.039)	−0.626 ^a (0.073)	−0.573 ^a (0.070)	−0.593 ^a (0.072)	−0.589 ^a (0.073)
IndMedian	0.360 ^a (0.051)	0.295 ^a (0.045)	0.328 ^a (0.048)	0.330 ^a (0.048)	0.081 (0.073)	0.042 (0.067)	0.051 (0.068)	0.069 (0.070)
N	4,669	4,669	4,669	4,669	1,896	1,896	1,896	1,896
Firms	901	901	901	901	678	678	678	678
R ²	0.27	0.29	0.28	0.28	0.28	0.31	0.30	0.29

Note: The table presents the estimation results of Equation (1) that relates the firm value (*Q*) to the crisis dummy (*Crisis*), financial conservatism dummy (*ZL*, *AZL* and *UL*) and firm-specific variables. The sample consists of listed non-utility and non-financial firms in selected African countries drawn from *Datastream* over the period 2003–2012. All variables used are defined in Table 1 and are winsorized at the lower and upper one percentiles. a, b, c and d indicate significance at the 0.01%, 1%, 5% and 10% levels, respectively.

manage exposure to international business risk, especially during periods of heightened economic uncertainty.

Similarly, we find that firms with foreign sales, in Panel B, were more affected by the GFC relative to firms without international sales. However, the negative effect of contractions in credit supply on the value of firms with foreign sales is less pronounced for financially conservative firms. These results are consistent with our second (H2) and third (H3) hypotheses. These findings suggest that even-though internationalization is beneficial for firms, it also inadvertently increases exposure to international market risks and is a channel through which economic shocks are transmitted across countries. However, as our results show, this negative effect diminishes with financial conservatism.

Finally, in Table 7, we examine whether the effects of the GFC on firm value vary with legal origin and whether financial conservatism moderates this effect. We estimate Equation (2) that includes a dummy for civil law countries and its interaction with the crisis and financial conservatism dummies (*Crisis*×*Civil*×*FC*). The coefficient of the *Crisis* dummy is consistently negative and in line with our prior results. We further find that the coefficient of the interaction term, *Crisis*×*Civil*, is positive and significant. This suggests that, unlike firms in common law countries, those in civil law countries were immune to the crisis since they are less integrated with the US and the UK where the crisis originated. The results are consistent with Figure 2, which shows that Kenya, Egypt, South Africa and Nigeria were more adversely affected by the crisis, while the effect is less pronounced for Ivory

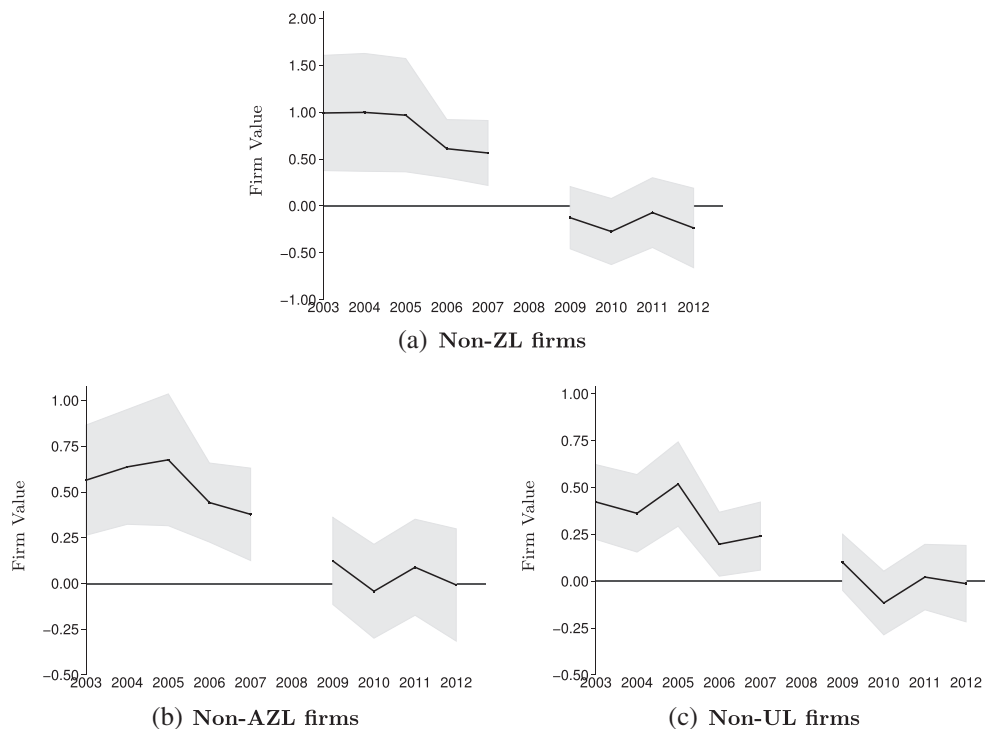


FIGURE 3 The changes in firm value of non-financially conservative relative to financially conservative firms. The figures plot the coefficients estimation Equation (1) with indicator variations for 2003–2007, 2009, 2010, 2011 and 2012 interacted with the dummy for non-financially conservative firms. The indicator for 2008 is omitted so that the coefficients from the regressions are interpreted as measuring the firm value are relative to that in 2008. The sample consists of listed non-utility and non-financial firms in selected African countries drawn from *Datastream* over the period 2003–2012. All variables used are defined in Table 1 and are winsorized at the lower and upper one percentiles

Coast, Tunisia and Morocco. This suggests that the effects of the GFC vary with legal system, and countries that are more tied to the US and the UK are more affected than others. Columns (2)–(4) further shows that the interaction terms of financial conservatism and the crisis dummy are consistently positive and significant, which supports Hypothesis 4 (H4). Columns (5)–(8) show similar results (for the period 2006–2009) which suggests that our findings are robust to changes in the window-period around the GFC. Overall, the results show that financial conservatism reduces the adverse effects of the credit supply shocks in common law countries while it enhances firm value in civil law countries.

In summary, our results confirm the adverse effect of the GFC on firm value as reported in the US by Duchin et al. (2010) and Almeida et al. (2011). We also show that the effect of the GFC on firm value varies with the level of leverage and international business exposure. Our further analyses, however, reveal that financially conservative firms and those without international exposure were less affected by the GFC than high-levered firms and those with international business exposure (the emerging market multinationals (EMMNEs)). More importantly, we find that financial conservatism reduces the adverse

effect of internationalization on firm value during the GFC. Hence, our results highlight the significant effect of financing activities on real decisions, which becomes more apparent during periods of heightened uncertainty, such as the GFC.

6 | ROBUSTNESS

In this section, we implement a battery of robustness tests. First, we re-estimate our baseline model of Equation (1) using MVE and TSR as the dependent variables. Table 8 summarizes the results for these alternative proxies of firm value.

Table 8 shows that, in all specifications, the coefficient of the crisis dummy (*Crisis*) is negative and significant, but the coefficients of the interaction terms between the financial conservatism dummy (ZL, AZL and UL, ZLL) and the crisis dummy (*Crisis*) are positive and significant. This finding is consistent with our central hypothesis (H1) and implies that firms in these countries were adversely affected by the 2008–09 credit supply shock despite being less integrated with developed economies, and that this effect is less pronounced for

TABLE 5 The effects of leverage policies on firm value

Variables	2003–2012		2006–2009	
	(1)	(2)	(3)	(4)
Constant	1.968 ^a (0.569)	2.005 ^a (0.588)	10.161 ^a (1.073)	9.949 ^a (1.091)
D2	−0.021 (0.038)	0.027 (0.050)	−0.095 (0.062)	−0.035 (0.081)
D3	−0.034 (0.043)	0.072 (0.058)	−0.168 ^c (0.072)	−0.065 (0.080)
D4	−0.089 ^d (0.048)	0.016 (0.062)	−0.180 ^c (0.083)	−0.107 (0.097)
Crisis		−0.159 ^c (0.067)		−0.341 ^a (0.079)
Crisis×D2		−0.090 (0.059)		−0.098 (0.083)
Crisis×D3		−0.202 ^b (0.062)		−0.197 ^c (0.095)
Crisis×D4		−0.199 ^b (0.063)		−0.145 ^d (0.086)
ROA	0.982 ^a (0.084)	0.945 ^a (0.081)	0.423 ^c (0.204)	0.285 (0.228)
SG	0.130 ^a (0.038)	0.131 ^a (0.038)	0.129 (0.089)	0.122 (0.087)
Cash	0.978 ^a (0.145)	0.961 ^a (0.150)	1.329 ^a (0.267)	1.331 ^a (0.276)
Size	−0.077 ^c (0.038)	−0.063 (0.038)	−0.604 ^a (0.071)	−0.561 ^a (0.073)
IndMedian	0.364 ^a (0.051)	0.350 ^a (0.049)	0.083 (0.073)	0.077 (0.071)
N	4,669	4,669	1,896	1,896
Firms	901	901	678	678
R ²	0.27	0.28	0.29	0.29

Note: The table presents the estimation results of Equation (1) that relates the firm value (Q) to the crisis dummy (*Crisis*), leverage policy dummies (D2, D3, and D4) and firm-specific variables. The sample consists of listed non-utility and non-financial firms in selected African countries drawn from *DataStream* over the periods 2003–2012 and 2006–2009. All variables used are defined in Table 1 and are winsorized at the lower and upper one percentiles. a, b, c and d indicate significance at the 0.01%, 1%, 5% and 10% levels, respectively.

conservative firms. Columns (1) and (2) show the robustness of our results to the use of alternative proxies of firm value and confirms that financial conservatism preserves or enhances firm value during the GFC.⁸

Next, in Columns (3)–(8) of Table 8, we replicate the main analyses using sub-samples for South Africa and

other countries. The aim is to ensure that our results are not mainly driven by one country (in this case South Africa). Columns (3)–(5) show that the coefficient of the crisis dummy (*Crisis*) is negative and significant, which suggests that the crisis had a similar adverse effect on firm value for firms in South Africa and other countries as shown in the previous sections. Columns (3)–(5) show that the interaction term, *Crisis*×*ZL*, is positive and significant, which suggests that financially conservative firms were less affected by the credit supply shock. Columns (6)–(8) show similar results for firms in other countries (excluding South Africa). Based on these results, we conclude that our baseline results are unlikely to be driven by one or a few countries in our sample.

We further conduct a robustness test where we re-estimate our baseline model across four industrial sectors, namely basic materials (BM), consumer goods and services (CG&S), health and technology (H&T) and industrial (IND) sectors. Consistent with our main results, Columns (9)–(12) show that the crisis had a significant adverse effect across all four sectors. Similarly, we find that financially conservative firms across all sectors, except for those in the H&T sectors, were less affected by the crisis. We attribute the somewhat puzzling result for the H&T sectors to binding financial constraints as untabulated results (available from the authors) show that these firms are smaller and younger than their non-conservative counterparts. A combination of the aforementioned characteristics in addition to operating in a very risky sector result to the adverse effects of the crisis outweighing the benefits of financial conservatism, thus leading to the negative overall coefficient (*Crisis* + *Crisis*×*ZL*). Based on this finding, we interpret the financial conservatism that we observe for firms in the H&T sector as externally imposed (due to binding credit constraints) rather than being deliberate or due to internal strategic choice.

Finally, we estimate a modified version of Equation (1) in which we replace the crisis dummy (*Crisis*) and the interaction term (*Crisis*×*FC_{it}*) with lagged corporate debt (*Debt_{ijt}* − 1). This enables us to directly examine how corporate debt affects firm value. In Figure 4, we plot the coefficient of lagged corporate debt for the cross-sectional yearly regressions of firm value while controlling for ROA, SG, cash and cash equivalent (Cash), PPE, size (Size) and the industrial median of market value (IndMedQ). The yearly plots of the coefficients show significant variation in the effect of corporate debt on firm value. Over the sample period, the negative effect of corporate debt on firm value peaks in 2006, which implies that symptoms of the crisis were visible earlier in Africa. This negative effect dissipates (Figure 4(a)) or is stable (Figure 4(b)) from 2006 onwards, which suggests

TABLE 6 Internationalization, financial conservatism and firm value

Panel A: Dual listing								
Variables	No (1)	Yes (2)	No (3)	Yes (4)	No (5)	Yes (6)	No (7)	Yes (8)
Constant	2.640 ^b (0.825)	1.867 ^c (0.815)	2.399 ^b (0.793)	1.996 ^c (0.780)	2.464 ^b (0.791)	1.769 ^c (0.783)	2.496 ^b (0.818)	1.530 ^d (0.816)
Crisis	−0.391 ^a (0.056)	−0.664 ^a (0.050)	−0.440 ^a (0.057)	−0.721 ^a (0.051)	−0.447 ^a (0.056)	−0.725 ^a (0.049)	−0.429 ^a (0.057)	−0.756 ^a (0.051)
Crisis×ZL			0.493 ^b (0.186)	1.196 ^a (0.105)				
Crisis×AZL					0.329 ^c (0.128)	0.524 ^a (0.143)		
Crisis×UL							0.128 (0.095)	0.414 ^a (0.094)
ROA	1.022 ^a (0.144)	0.923 ^a (0.102)	0.996 ^a (0.139)	0.830 ^a (0.087)	0.987 ^a (0.142)	0.849 ^a (0.094)	1.011 ^a (0.143)	0.903 ^a (0.094)
SG	0.102 ^c (0.052)	0.176 ^b (0.058)	0.093 ^d (0.049)	0.165 ^b (0.053)	0.098 ^c (0.050)	0.182 ^b (0.056)	0.103 ^c (0.051)	0.186 ^b (0.056)
Cash	0.746 ^a (0.190)	1.381 ^a (0.209)	0.678 ^a (0.201)	1.351 ^a (0.207)	0.715 ^a (0.198)	1.373 ^a (0.212)	0.754 ^a (0.193)	1.368 ^a (0.214)
Size	−0.098 ^d (0.057)	−0.076 (0.051)	−0.075 (0.054)	−0.068 (0.049)	−0.082 (0.054)	−0.062 (0.049)	−0.086 (0.056)	−0.046 (0.052)
IndMedian	0.220 ^a (0.056)	0.606 ^a (0.084)	0.184 ^a (0.051)	0.475 ^a (0.086)	0.200 ^a (0.054)	0.555 ^a (0.084)	0.210 ^a (0.055)	0.542 ^a (0.081)
N	2,371	2,298	2,371	2,298	2,371	2,298	2,371	2,298
Firms	515	386	515	386	515	386	515	386
R ²	0.23	0.33	0.25	0.37	0.24	0.35	0.23	0.35
Panel B: Foreign sales								
Variables	No (1)	Yes (2)	No (3)	Yes (4)	No (5)	Yes (6)	No (7)	Yes (8)
Constant	2.523 ^a (0.655)	1.319 (0.932)	2.213 ^a (0.633)	1.405 (0.915)	2.267 ^a (0.633)	1.423 (0.916)	2.319 ^a (0.658)	1.744 ^d (0.911)
Crisis	−0.387 ^a (0.048)	−0.794 ^a (0.070)	−0.454 ^a (0.048)	−0.838 ^a (0.072)	−0.452 ^a (0.048)	−0.858 ^a (0.062)	−0.429 ^a (0.051)	−0.872 ^a (0.063)
Crisis×ZL			0.639 ^a (0.179)	0.855 ^a (0.209)				
Crisis×AZL					0.321 ^b (0.102)	0.773 ^b (0.251)		
Crisis×UL							0.120 (0.075)	0.681 ^a (0.167)
ROA	0.843 ^a (0.101)	1.277 ^a (0.130)	0.840 ^a (0.099)	1.177 ^a (0.115)	0.822 ^a (0.101)	1.168 ^a (0.115)	0.841 ^a (0.102)	1.138 ^a (0.111)
SG	0.087 ^c (0.042)	0.264 ^a (0.066)	0.075 ^d (0.040)	0.253 ^a (0.063)	0.085 ^c (0.041)	0.259 ^a (0.063)	0.088 ^c (0.042)	0.267 ^a (0.063)

(Continues)

TABLE 6 (Continued)

Panel B: Foreign sales								
Variables	No (1)	Yes (2)	No (3)	Yes (4)	No (5)	Yes (6)	No (7)	Yes (8)
Cash	0.675 ^a (0.157)	1.480 ^a (0.278)	0.605 ^a (0.172)	1.412 ^a (0.268)	0.662 ^a (0.163)	1.428 ^a (0.268)	0.691 ^a (0.158)	1.434 ^a (0.277)
Size	−0.099 ^c (0.048)	−0.054 (0.055)	−0.067 (0.046)	−0.058 (0.054)	−0.076 ^d (0.046)	−0.058 (0.054)	−0.082 ^d (0.048)	−0.069 (0.054)
IndMedian	0.264 ^a (0.046)	0.752 ^a (0.167)	0.202 ^a (0.036)	0.761 ^a (0.166)	0.243 ^a (0.040)	0.754 ^a (0.161)	0.256 ^a (0.043)	0.668 ^a (0.164)
N	2,720	1,949	2,720	1,949	2,720	1,949	2,720	1,949
Firms	675	354	675	354	675	354	675	354
R ²	0.23	0.35	0.26	0.36	0.24	0.37	0.23	0.38

Note: The table presents the estimation results of Equation (1) that relates the firm value (Q) to the crisis dummy (*Crisis*), financial conservatism dummy (ZL, AZL and UL) and firm-specific variables. Figure (a) includes all firms while Figure (b) is for the restricted sample excluding firms not listed or headquartered in Africa. The sample consists of listed non-utility and non-financial firms in selected African countries drawn from *Datastream* over the period 2003–2012. All variables used are defined in Table 1 and are winsorized at the lower and upper one percentiles. a, b, c and d indicate significance at the 0.01%, 1%, 5% and 10% levels, respectively.

that African firms recovered much earlier than those in developed countries. The stable coefficient for the restricted sample of only firms that are listed in Africa over the period 2006–2010, in Figure 4(b), further suggests that firms with operations outside Africa were more affected than local or domestic firms. These results, as shown previously in Table 6, indicate that cross listing is one of the channels through which credit supply shocks are transmitted across capital markets. Overall, our robustness analyses confirm that financing policies have a significant effect on firm value as is consistent with our main results.

7 | SUMMARY DISCUSSION AND CONCLUSION

Motivated by the ongoing debate on whether managers leave money on the table by adopting conservative financing policies, we use a large sample of African firms and the GFC as a *quasi-natural experiment* to study the effects of financial conservatism on firm value. Specifically, we examine the effects of the GFC on firm value and how this is moderated by financial conservatism. We then examine how the effects of the GFC and the moderating role of financial conservatism vary across local and international firms and across countries with different legal systems.

Our study makes several contributions to theory and practice. First, we find robust and significant adverse effects of the GFC on the value of publicly listed firms in emerging markets that are less integrated with developed economies. Our further analyses show that, relative to

firms in civil law countries, the adverse effects of the GFC were more pronounced for firms in common law countries that have stronger ties to the US and the UK where the financial crisis originated. This provides new evidence on the far-reaching effects of the 2008–2009 contraction in credit supply and how ties in legal systems aid the transmission of economic shocks across countries.

Second, we find that the adverse effect of the GFC is, however, less pronounced for financially conservative firms, which suggests that adopting conservative financing policies preserve firm value during crises periods. This is contrary to the mainstream literature, which shows that using debt financing is beneficial as it increases the interest tax shield, lowers financing costs and mitigates the extraction of private benefits by managers. Instead, our results suggest that such benefits associated with debt financing may not sufficiently outweigh those arising from financial conservatism (in the form of financial flexibility), particularly during periods of marked contractions in credit supply.

Third, we show that strategic adoption of financial conservatism is one way of effectively managing risk, especially during economic downturns and for firms with exposure to international business risk (firms with foreign sales or cross-listings). Our results suggest that financial conservatism increases financial flexibility, thereby immunizing firms against credit supply shocks in both local and international capital markets. By showing that the effects of the GFC vary with financing policies, we also provide further and more direct evidence on the contentious nexus between financing and real corporate activities. This also serves as a forewarning of the potential adverse effects of over-leveraging and the need to

TABLE 7 Legal origin, financial conservatism and firm value

Variables	2003–2012				2006–2009			
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)
Constant	2.344 ^a (0.591)	2.182 ^a (0.569)	2.189 ^a (0.570)	2.087 ^a (0.585)	10.623 ^a (1.076)	10.197 ^a (1.045)	10.356 ^a (1.063)	10.280 ^a (1.067)
Crisis	−0.544 ^a (0.038)	−0.574 ^a (0.039)	−0.582 ^a (0.038)	−0.597 ^a (0.040)	−0.473 ^a (0.040)	−0.504 ^a (0.039)	−0.506 ^a (0.041)	−0.509 ^a (0.043)
Crisis×civil	0.909 ^a (0.128)	0.370 ^a (0.108)	0.374 ^a (0.108)	0.387 ^a (0.108)	0.832 ^a (0.114)	0.412 ^a (0.093)	0.414 ^a (0.093)	0.417 ^a (0.093)
Crisis×ZL		0.586 ^b (0.179)				0.412 ^c (0.175)		
Crisis×civil×ZL		0.259 (0.244)				0.233 (0.229)		
Crisis×AZL			0.322 ^b (0.100)				0.214 ^c (0.108)	
Crisis×civil×AZL			0.515 ^b (0.191)				0.431 ^c (0.185)	
Crisis×UL				0.224 ^b (0.070)				0.121 (0.076)
Crisis×civil×UL				0.618 ^a (0.178)				0.521 ^b (0.168)
ROA	0.975 ^a (0.084)	0.933 ^a (0.078)	0.932 ^a (0.081)	0.957 ^a (0.080)	0.411 ^c (0.201)	0.309 (0.204)	0.336 (0.209)	0.350 (0.214)
SG	0.126 ^a (0.038)	0.121 ^a (0.037)	0.129 ^a (0.037)	0.133 ^a (0.038)	0.118 (0.088)	0.105 (0.086)	0.125 (0.088)	0.130 (0.089)
Cash	1.006 ^a (0.144)	0.959 ^a (0.150)	0.992 ^a (0.148)	1.016 ^a (0.148)	1.306 ^a (0.266)	1.298 ^a (0.264)	1.329 ^a (0.266)	1.338 ^a (0.265)
Size	−0.082 ^c (0.039)	−0.065 ^d (0.037)	−0.067 ^d (0.038)	−0.060 (0.039)	−0.611 ^a (0.073)	−0.578 ^a (0.070)	−0.589 ^a (0.072)	−0.585 ^a (0.072)
IndMedian	0.322 ^a (0.046)	0.281 ^a (0.045)	0.298 ^a (0.046)	0.294 ^a (0.045)	0.066 (0.066)	0.040 (0.067)	0.043 (0.067)	0.052 (0.067)
N	4,669	4,669	4,669	4,669	1,896	1,896	1,896	1,896
Firms	901	901	901	901	678	678	678	678
R ²	0.28	0.30	0.29	0.29	0.31	0.32	0.32	0.32

Note: The table presents the estimation results of Equation (1) that relates the firm value (*Q*) to the crisis dummy (*Crisis*), financial conservatism dummy (*ZL*, *AZL* and *UL*), civil law dummy (*Civil*), and firm-specific variables. The sample consists of listed non-utility and non-financial firms in selected African countries drawn from *Datastream* over the periods 2003–2012 and 2006–2009. All variables used are defined in Table 1 and are winsorized at the lower and upper one percentiles. a, b, c and d indicate significance at the 0.01%, 1%, 5% and 10% levels, respectively.

develop robust risk management strategies, especially given the recent surge in corporate debt (Lund, Woetzel, Windhagen, Dobbs, & Goldshtein, 2018; Machokoto et al., 2020) and predictions of another eminent crisis as firms appear to be taking investment risks similar to those that led to the 2008–2009 GFC (see Duffie, 2019; Syriopoulos, Makram, & Boubaker, 2015).

Finally, contrary to prior literature on the bonding hypothesis (e.g., Areneke & Kimani, 2019; Charitou & Louca, 2009; Coffee, 2002; Lel & Miller, 2008) which documents significant positive spillover effects on firm value, we find that bonding is not always beneficial, especially during crises periods as it increases contagion risks. Our results show that firms which cross-list to gain access to

TABLE 8 Alternative sub-samples and definitions of firm value

Variables	All firms			South Africa			Other countries				Sectorial analysis			
	MVE	TSR	Q	MVE	TSR	Q	MVE	TSR	BM	Cg&S	H&T	IND		
	(1)	(2)	(3)	(4)	(5)	(6)	(7)	(8)	Q	(9)	(10)	(11)	(12)	
Constant	4.877 ^a (1.137)	4.991 ^a (0.698)	1.687 ^b (0.585)	3.792 ^b (1.190)	5.449 ^a (0.774)	1.667 (1.773)	4.521 (3.713)	5.182 ^b (1.675)	3.943 ^a (1.073)	-2.291 (1.450)	1.914 (1.396)	4.701 ^a (0.662)		
Crisis	-0.802 ^a (0.088)	-0.249 ^a (0.037)	-0.713 ^a (0.044)	-0.622 ^a (0.138)	-0.228 ^a (0.042)	-0.357 ^a (0.097)	-0.771 ^a (0.160)	-0.326 ^a (0.076)	-0.689 ^a (0.082)	-0.437 ^a (0.087)	-0.530 ^a (0.140)	-0.558 ^a (0.051)		
Crisis×ZL	1.136 ^a (0.228)	0.328 ^a (0.048)	0.384 ^d (0.197)	0.424 ^d (0.221)	0.258 ^a (0.044)	0.981 ^a (0.214)	1.543 ^a (0.329)	0.248 ^c (0.096)	1.697 ^a (0.236)	0.338 ^d (0.174)	-0.316 ^b (0.115)	0.527 ^a (0.135)		
ROA	1.627 ^a (0.177)	-0.489 ^a (0.098)	0.890 ^a (0.078)	1.610 ^a (0.184)	-0.423 ^a (0.098)	1.216 ^b (0.417)	1.400 (0.900)	-0.462 (0.339)	0.417 ^a (0.053)	2.238 ^a (0.331)	1.518 ^a (0.412)	0.994 ^a (0.214)		
SG	0.284 ^a (0.073)	0.120 ^c (0.050)	0.119 ^b (0.039)	0.302 ^a (0.077)	0.099 ^d (0.054)	0.214 ^d (0.115)	0.240 (0.223)	0.288 ^c (0.123)	0.207 ^b (0.068)	0.031 (0.123)	-0.148 (0.130)	0.152 ^b (0.050)		
Cash	1.599 ^a (0.308)	0.144 (0.134)	1.108 ^a (0.154)	1.910 ^a (0.326)	0.219 (0.150)	0.177 (0.353)	0.545 (0.724)	-0.061 (0.305)	1.365 ^a (0.198)	0.646 ^d (0.345)	0.562 ^d (0.299)	0.904 ^a (0.196)		
Size	-0.190 ^c (0.076)	-0.303 ^a (0.047)	-0.064 ^d (0.037)	-0.151 ^d (0.078)	-0.328 ^a (0.051)	-0.016 (0.127)	-0.172 (0.266)	-0.341 ^b (0.121)	-0.141 ^c (0.067)	0.240 ^c (0.097)	-0.057 (0.112)	-0.231 ^a (0.044)		
IndMedian	0.222 ^a (0.041)	-0.017 (0.046)	0.603 ^a (0.100)	0.468 ^a (0.098)	-0.040 (0.055)	0.092 ^c (0.045)	0.074 ^d (0.043)	-0.266 ^a (0.067)	0.153 ^b (0.059)	0.143 (0.096)	0.237 (0.239)	0.149 ^c (0.064)		
N	4,669	4,669	3,647	3,647	3,647	1,022	1,022	1,022	1,367	1,333	377	1,592		
Firms	901	901	647	647	647	254	254	254	261	259	71	310		
R ²	0.26	0.15	0.36	0.31	0.20	0.17	0.13	0.14	0.41	0.30	0.37	0.40		

Note: The table presents the estimation results of Equation (1) that relates the firm value to the crisis dummy (*Crisis*), financial conservatism dummy (*ZL*) and firm-specific variables. The three alternative measures of firm value are firm value (*Q*), the market value-to-equity (*MVE*) and total shareholder return (*TSR*). Columns (1) and (2) includes all firms, while Columns (3)–(5) is for firms in South Africa and Columns (6)–(8) is for firms from other countries (excluding South Africa). Column (9) is for firms in the Basic Materials (BM) sector. Column (10) is for firms in the Consumer Goods and Services (CG&S) sectors. Column (11) is for firms in the Health and Technology (H&T) sectors. Column (12) is for firms in the Industrial (IND) sector. The sample consists of listed non-utility and non-financial firms in selected African countries drawn from *Datastream* over the period 2003–2012. All variables used are defined in Table 1 and are win-sorted at the lower and upper one percentiles. a, b, c and d indicate significance at the 0.01%, 1%, 5% and 10% levels, respectively.

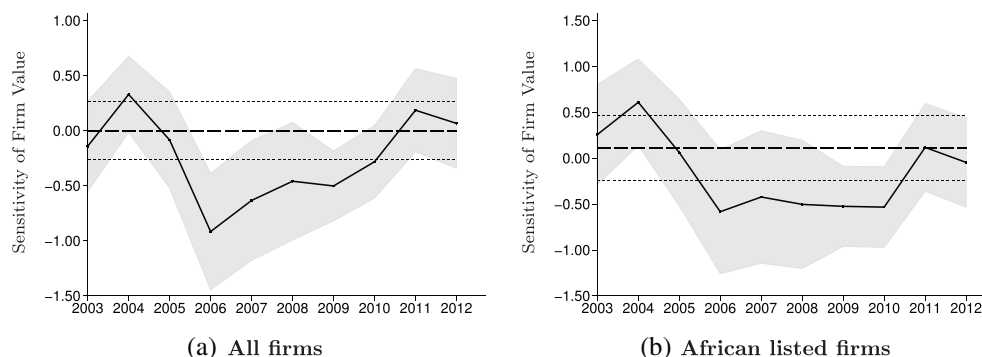


FIGURE 4 The effects of leverage on firm value around the crisis. The figure plots the coefficient of leverage (sensitivity of firm value) in the cross-sectional yearly regressions of firm value, controlling for return on assets (ROA), sales growth (SG), cash and cash equivalent (Cash), PPE, size (Size), and the industrial median of market value (IndMedQ). The sample consists of listed non-utility and non-financial firms in selected African countries drawn from *Datastream* over the period 2003–2012. All variables used are defined in Table 1 and are winsorized at the lower and upper one percentiles

external finance and benefit from better governance systems (following on the bonding hypothesis of Coffee (2002)), fare worse than domestic firms during the GFC, except if they are financially conservative. Overall, our findings yield strong support for the moderating role of financial conservatism on firm value during the 2008–2009 contractions in credit supply.

ACKNOWLEDGEMENTS

We wish to thank an anonymous referee and The Editor for several insightful and constructive comments that help improve our paper. We also acknowledge the helpful input of Boulis Ibrahim, Nadeem Aftab, Thaana Ghalia and seminar participants at the Research Discussion Group Meetings (University of Northampton).

DATA AVAILABILITY STATEMENT

The data that support the findings of this study are available from *Thompson Reuters Datastream*. Restrictions apply to the availability of these data, which were used under license for this study.

ORCID

Michael Machokoto  <https://orcid.org/0000-0001-8903-0019>

Geofry Areneke  <https://orcid.org/0000-0001-6075-2747>

ENDNOTES

¹ Several studies report contradictory results on the effects of the GFC on corporate decisions. For example, Duchin, Ozbas, and Sensoy (2010) and Almeida, Campello, and Weisbach (2011) find that the crisis had a significant adverse effect on corporate investment. However, Kahle and Stulz (2013) find no significant

differences between non-bank-dependent and bank-dependent firms of the effect of the crisis on debt issuance and capital expenditure.

² The channels through which financing activities affect real decisions are subject to debate (see Chava & Roberts, 2008; Stein, 2003). Since the credit supply shock originated in developed countries (the US and the UK), it should be orthogonal to African firms and their operating environment. This ought to reduce concerns of compounding effects (Popov & Rocholl, 2018; Puri, Rocholl, & Steffen, 2011).

³ For studies that examine the determinants of zero-leverage see Devos, Dhillon, Jagannathan, and Krishnamurthy (2012), Bessler et al. (2013), Dang (2013), and El Ghouli, Guedhami, Kwok, and Zheng (2018).

⁴ This addresses the concentration of studies on the effects of the GFC in the US (Almeida & Campello, 2007; Chava & Purnanandam, 2011; Kahle & Stulz, 2013).

⁵ Similarly, Graham and Leary (2011) also conclude that covenant violations during the GFC did not significantly limit access to further credit as lenders were more willing to renegotiate.

⁶ The choice of firm-specific factors is informed by the literature (e.g., Fama & French, 1998; Gamba & Triantis, 2008; Kim, Park, & Suh, 2018).

⁷ See Ojah and Pillay (2009), Agyei-Boapeah and Machokoto (2018) and Machokoto, Areneke, and Ibrahim (2020).

⁸ In unreported results, we include a set of macroeconomic control variables that may affect firm value and find that our main conclusions do not change. This should allay concerns of other omitted variables such as macroeconomic factors biasing our inferences.

REFERENCES

- Aggarwal, R., Erel, I., Stulz, R., & Williamson, R. (2009). Differences in governance practices between U.S. and foreign firms: Measurement, causes, and consequences. *The Review of Financial Studies*, 22, 3131–3169.

- Aggei-Boapeah, H., & Machokoto, M. (2018). Allocation of internally generated corporate cash flow in Africa. *Journal of Accounting in Emerging Economies*, 8, 495–513.
- Allen, F., & Giovannetti, G. (2011). The effects of the financial crisis on sub-Saharan Africa. *Review of Development Finance*, 1, 1–27.
- Almeida, H., & Campello, M. (2007). Financial constraints, asset tangibility, and corporate investment. *Review of Financial Studies*, 20, 1429–1460.
- Almeida, H., Campello, M., & Weisbach, M. S. (2011). Corporate financial and investment policies when future financing is not frictionless. *Journal of Corporate Finance*, 17, 675–693.
- Almeida, H., Murillo, C., Bruno, L., & Scott, W. (2012). Corporate debt maturity and the real effects of the 2007 credit crisis. *Critical Finance Review*, 1, 3–58.
- Anginer, D., & Demirguc-Kunt, A. (2014). Has the global banking system become more fragile over time? *Journal of Financial Stability*, 13, 202–213.
- Angrist, J. D., & Pischke, J.-S. (2008). *Mostly harmless econometrics: An Empiricist's companion*. Princeton University Press Google-Books-ID: ztXL21Xd8v8C. <https://press.princeton.edu/books/paperback/9780691120355/mostly-harmless-econometrics>
- Areneke, G., & Kimani, D. (2019). Value relevance of multinational directorship and cross-listing on MNEs national governance disclosure practices in sub-Saharan Africa: Evidence from Nigeria. *Journal of World Business*, 54, 285–306.
- Baek, J.-S., Kang, J.-K., & Suh Park, K. (2004). Corporate governance and firm value: Evidence from the Korean financial crisis. *Journal of Financial Economics*, 71, 265–313.
- Berman, N., & Martin, P. (2012). The vulnerability of sub-Saharan Africa to financial crises: The case of trade. *IMF Economic Review*, 60, 329–364.
- Bessler, W., Drobetz, W., Haller, R., & Meier, I. (2013). The international zero-leverage phenomenon. *Journal of Corporate Finance*, 23, 196–221.
- Bigelli, M., Martín-Ugedo, J. F., & Sánchez-Vidal, F. J. (2014). Financial conservatism of private firms. *Journal of Business Research*, 67, 2419–2427.
- Campello, M., Graham, J. R., & Harvey, C. R. (2010). The real effects of financial constraints: Evidence from a financial crisis. *Journal of Financial Economics*, 97, 470–487.
- Caporale, G. M., Lodh, S., & Nandy, M. (2018). How has the global financial crisis affected syndicated loan terms in emerging markets? Evidence from China. *International Journal of Finance & Economics*, 23, 478–491.
- Chari, V. V., Christiano, L. J., & Kehoe, P. J. (2008). *Facts and myths about the financial crisis of 2008, working paper 666*. Federal Reserve Bank of Minneapolis: Minneapolis, MN.
- Charitou, A., & Louca, C. (2009). Cross-listing and operating performance: Evidence from exchange-listed American depository receipts. *Journal of Business Finance & Accounting*, 36, 99–129.
- Chava, S., & Purnanandam, A. (2011). The effect of banking crisis on Bank-dependent borrowers. *Journal of Financial Economics*, 99, 116–135.
- Chava, S., & Roberts, M. R. (2008). How does financing impact investment? The role of debt covenants. *The Journal of Finance*, 63, 2085–2121.
- Chevalier, J., & Scharfstein, D. (1996). Capital-market imperfections and countercyclical markups: Theory and evidence. *American Economic Review*, 86, 703–725.
- Choudhry, T., Hassan, S. S., & Shabi, S. (2020). U.S. economic uncertainty, EU business cycles, and the global financial crisis. *International Journal of Finance & Economics*, 25(1), 28–42. <http://dx.doi.org/10.1002/ijfe.1726>.
- Coffee, J. C. (2002). Racing towards the top?: The impact of cross-listings and stock market competition on international corporate governance. *Columbia Law Review*, 102, 1757–1831.
- Dang, V. A. (2013). An empirical analysis of zero-leverage firms: New evidence from the UK. *International Review of Financial Analysis*, 30, 189–202.
- Dawood, M., Horsewood, N., & Strobel, F. (2017). Predicting sovereign debt crises: An early warning system approach. *Journal of Financial Stability*, 28, 16–28.
- Devos, E., Dhillon, U., Jagannathan, M., & Krishnamurthy, S. (2012). Why are firms unlevered? *Journal of Corporate Finance*, 18, 664–682.
- Duchin, R., Ozbas, O., & Sensoy, B. A. (2010). Costly external finance, corporate investment, and the subprime mortgage credit crisis. *Journal of Financial Economics*, 97, 418–435.
- Duffie, D. (2019). Prone to fail: The pre-crisis financial system. *Journal of Economic Perspectives*, 33, 81–106.
- Duygun, M., Isaksson, A., & Kaulich, F. (2016). Did the global financial crisis hit Africa? Insights from a multi-country firm level survey. *Economic Modelling*, 58, 308–316.
- El Ghouli, S., Guedhami, O., Kwok, C., & Zheng, X. (2018). Zero-Leverage Puzzle: An International Comparison. *Review of Finance*, 22, 1063–1120.
- Enikolopov, R., Petrova, M., & Stepanov, S. (2014). Firm value in crisis: Effects of firm-level transparency and country-level institutions. *Journal of Banking & Finance*, 46, 72–84.
- Fainshmidt, S., Judge, W. Q., Aguilera, R. V., & Smith, A. (2018). Varieties of institutional systems: A contextual taxonomy of understudied countries. *Journal of World Business*, 53, 307–322.
- Fama, E. F., & French, K. R. (1998). Taxes, financing decisions, and firm value. *The Journal of Finance*, 53, 819–843.
- Gamba, A., & Triantis, A. (2008). The value of financial flexibility. *The Journal of Finance*, 63, 2263–2296.
- Graham, J. R., & Harvey, C. R. (2001). The theory and practice of corporate finance: Evidence from the field. *Journal of Financial Economics*, 60, 187–243.
- Graham, J. R., & Leary, M. T. (2011). A review of empirical capital structure research and directions for the future. *Annual Review of Financial Economics*, 3, 309–345.
- Gupta, K., Krishnamurti, C., & Tourani-Rad, A. (2013). Is corporate governance relevant during the financial crisis? *Journal of International Financial Markets, Institutions and Money*, 23, 85–110.
- Holmstrom, B., & Tirole, J. (1997). Financial intermediation, loanable funds, and the real sector. *The Quarterly Journal of Economics*, 112, 663–691.
- Hsu, W.-T., Chen, H.-L., & Cheng, C.-Y. (2013). Internationalization and firm performance of SMEs: The moderating effects of CEO attributes. *Journal of World Business*, 48, 1–12.
- Juasrikul, S., Sahaym, A., Yim, H. (S.), & Liu, R. L. (2018). Do cross-border alliances with MNEs from developed economies create firm value for MNEs from emerging economies? *Journal of Business Research*, 93, 98–110.
- Kahle, K. M., & Stulz, R. M. (2013). Access to capital, investment, and the financial crisis. *Journal of Financial Economics*, 110, 280–299.

- Kim, S., Park, S. H., & Suh, J. (2018). A J-shaped cross-sectional relation between dividends and firm value. *Journal of Corporate Finance*, 48, 857–877. <http://dx.doi.org/10.1016/j.jcorpfin.2016.09.010>.
- La Porta, R., Lopez-De-Silanes, F., & Shleifer, A. (2006). What works in securities Laws? *The Journal of Finance*, 61, 1–32.
- La Porta, R., Lopez-De-Silanes, F., Shleifer, A., & Vishny, R. W. (1997). Legal determinants of external finance. *The Journal of Finance*, 52, 1131–1150.
- Lel, U., & Miller, D. P. (2008). International cross-listing, firm performance, and top management turnover: A test of the bonding hypothesis. *The Journal of Finance*, 63, 1897–1937.
- Lemmon, M., & Roberts, M. R. (2010). The response of corporate financing and investment to changes in the supply of credit. *Journal of Financial and Quantitative Analysis*, 45, 555–587.
- Lund, S., Woetzel, J., Windhagen, E., Dobbs, R., & Goldshtein, D. (2018). *Rising corporate debt: Peril or promise?, discussion paper*. New York, NY: McKinsey Global Institute, McKinsey&Company.
- Machokoto, M., Areneke, G., & Ibrahim, B. M. (2020). Rising corporate debt and value relevance of supply-side factors in South Africa. *Journal of Business Research*, 109, 26–37.
- Ojah, K., & Pillay, K. (2009). Debt markets and corporate debt structure in an emerging market: The south African example. *Economic Modelling*, 26, 1215–1227.
- Popov, A., & Rocholl, J. (2018). Do credit shocks affect labor demand? Evidence for employment and wages during the financial crisis. *Journal of Financial Intermediation*, 36, 16–27.
- Puri, M., Rocholl, J., & Steffen, S. (2011). Global retail lending in the aftermath of the US financial crisis: Distinguishing between supply and demand effects. *Journal of Financial Economics*, 100, 556–578.
- Rehman, A., Chaudhry, S. M., & Hussain, S. M. (2019). Transmission of a global financial crisis shock to an emerging economy. *International Journal of Finance & Economics*, 24, 740–760.
- Stein, J. C. (2003). Agency, information and corporate investment. In M. Harris, R. M. Stulz, & G. M. Constantinides (Eds.), *Handbook of the economics of finance: Part a of corporate finance* (Vol. 1, pp. 111–165). Amsterdam, North-Holland: Elsevier.
- Syriopoulos, T., Makram, B., & Boubaker, A. (2015). Stock market volatility spillovers and portfolio hedging: BRICS and the financial crisis. *International Review of Financial Analysis*, 39, 7–18.
- Udenio, M., Hoberg, K., & Fransoo, J. C. (2018). Inventory agility upon demand shocks: Empirical evidence from the financial crisis. *Journal of Operations Management*, 62, 16–43.
- Zouaghi, F., Sánchez, M., & Martinez, M. G. (2018). Did the global financial crisis impact firms' innovation performance? The role of internal and external knowledge capabilities in high and low tech industries. *Technological Forecasting and Social Change*, 132, 92–104.

How to cite this article: Machokoto M, Areneke G, Nyangara D. Financial conservatism, firm value and international business risk: Evidence from emerging economies around the global financial crisis. *Int J Fin Econ*. 2020;1–19. <https://doi.org/10.1002/ijfe.2032>